European Establishment Statistics Workshop

12 – 14 september 2011 - Neuchâtel

Sampling coordination of business surveys conducted by INSEE

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Random numbers

Very often national statistics agencies use the permanent random numbers technique for the sampling coordination of business surveys.

Each unit k of the population (included the new units) is independently assigned a number ω_k , selected according to the uniform distribution in the interval [0,1 [.

Poisson sampling: selection of the units k whose numbers belong to the interval [d, d + π_k [, π_k = probability of inclusion of the unit k

Simple random sampling (SRS): selection of the n units with the lowest numbers ω_k superior to d.

Constant shift method: J panels (j = 1 ...J)

Starting point for panel j at the date a: $d_{j,a} = d_{j,1} + (a-1)c$, $a \ge 1$

Definition of a coordination function

Coordination function g = measurable application from [0,1[to [0,1[which preserves uniform probability :

if P is the uniform probability on [0,1[, then the image probability

$$P^g = P$$
.

 \rightarrow for any interval I = [a,b [included in [0,1[:

$$P(g^{-1}(I)) = P^{g}(I) = P(I) = b - a$$

Selection of the units

For each unit k in the sampling frame:

- a permanent random number ω_k ,
- a coordination function $g_{k,t}$ that changes at each sampling t = 1, 2, ...

1. Poisson sampling

Selection of the units k such that $g_k(\omega_k) \in [0, \pi_k[$ where π_k = probability of inclusion of the unit k

$$P(k \in S) = P(g_k(\omega_k) \in [0, \pi_k[) = P^{g_k}([0, \pi_k[) = P([0, \pi_k[) = \pi_k[) = \pi_k[) = \pi_k[) = P([0, \pi_k[) = \pi_k[) = \pi_k[) = \pi_k[] = \pi_k[]$$

and the drawings are independent.

Selection of the units

2. SSRS

Within a stratum, selection of the n units k associated with the n smallest numbers $g_k(\omega_k)$.

Since

- P^g_k = uniform probability P sur [0,1[for each k
- and the n numbers (ω_k) are drawn independently from P
- \rightarrow the n numbers $g_k(\omega_k)$ are drawn independently from P
- \rightarrow the n smallest numbers $g_k(\omega_k)$ give a simple random sample of size n in the stratum.

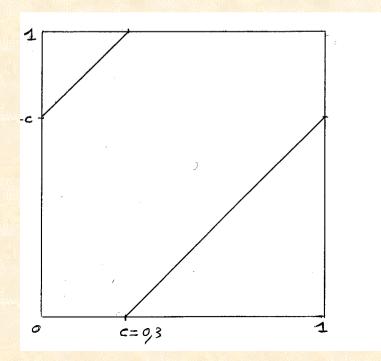
Example: the constant shift method

Let $d_1 = 0$ and $d_2 = c$. We define the coordination functions:

$$g_{k,1}(\omega_k) = \omega_k$$
 $g_{k,2}(\omega_k) = \omega_k - c \pmod{1}$

Then:

$$\begin{aligned} k \in S_1 & \Leftrightarrow \omega_k \in \left[0, \pi_{k,1}\right[\iff g_{k,1}(\omega_k) \in \left[0, \pi_{k,1}\right[\\ k \in S_2 & \Leftrightarrow \omega_k \in \left[c, c + \pi_{k,2}\right[\iff g_{k,2}(\omega_k) \in \left[0, \pi_{k,2}\right[\\ \end{aligned} \right] \end{aligned}$$



A step by step procedure reflecting response burdens

 $\boldsymbol{\omega} = (\dots \omega_k \dots) = \text{vector of random numbers given to the population units.}$

 $I_{k,t}(\omega)$ = indicator function, equal to 1 if the values in ω lead to select the unit k in the sampling t, and 0 otherwise :

$$k \in S_t \Leftrightarrow I_{k,t}(\boldsymbol{\omega}) = 1$$

 $\gamma_{k,t}$ = response burden of a questioned enterprise k at survey t

Effective burden = random variable $\gamma_{k,t}(\omega) = \gamma_{k,t} I_{k,t}(\omega)$

Cumulative burden for unit k: $\Gamma_{k,t}(\omega) = \sum_{v < t} \gamma_{k,u} I_{k,u}(\omega)$

Principle: to define the coordination functions $g_{k,t}$ for the selection of sample S_t using $\Gamma_{k,t-1}$:

$$\Gamma_{k,t-1}(\boldsymbol{\omega_1}) < \Gamma_{k,t-1}(\boldsymbol{\omega_2}) \Rightarrow g_{k,t}(\omega_{k,1}) < g_{k,t}(\omega_{k,2})$$

Two difficulties

1. to substitute for $\Gamma_{k,t}(\boldsymbol{\omega})$ a function of ω_k only, denoted $\Gamma'_{k,t}(\omega_k)$, that closely approximates $\Gamma_{k,t}(\boldsymbol{\omega})$.

Poisson sampling: $I_{k,t}(\omega)$ depends only on ω_k (indicator function of an interval of length π_k), and can be denoted $I_{k,t}(\omega_k)$ $\to \Gamma_{k,t}(\omega)$ depends only on ω_k , and can be denoted $\Gamma'_{k,t}(\omega_k)$.

<u>SSRS</u>: $I_{k,t}(\omega)$ depends on <u>all coordinates</u> of the vector ω , but "primarily" on coordinate ω_k : if we select the n units among N with the n smallest values ω_j , it will be equal to 1 for values of ω_k near to 0, regardless of the values of the other coordinates \rightarrow we will be able to replace $I_{k,t}(\omega)$ with an approximation $I'_{k,t}(\omega_k)$, and therefore to replace $\Gamma_{k,t}(\omega)$ with an approximation $\Gamma'_{k,t}(\omega_k)$.

2. to define the coordination function $g_{k,t}$ such that :

$$\Gamma'_{k,t-1}(\omega_{k,1}) < \Gamma'_{k,t-1}(\omega_{k,2}) \Longrightarrow g_{k,t}(\omega_{k,1}) < g_{k,t}(\omega_{k,2})$$

Construction of a coordination function

 $C_{k,t}$ (ω_k) = criterion such that the smaller is the criterion, the larger is the probability of selection for unit k at sampling t.

We drop the subscripts k and t. C is supposed to be a bounded measurable function:

$$\omega \in [0,1[\rightarrow C(\omega) \in IR]$$

Idea: to associate to this criterion a coordination function g_C such that:

$$C(\omega_1) < C(\omega_2) \Rightarrow g_C(\omega_1) < g_C(\omega_2)$$
 (1)

 P^{C} = image probability of P under C, F_{C} the distribution function of C.

The coordination function is built from $G_C = F_C(C)$:

$$G_{C}(\omega) = P^{C}(]-\infty, C(\omega)[) = P(C^{-1}]-\infty, C(\omega)[) = P(u|C(u) < C(\omega))$$

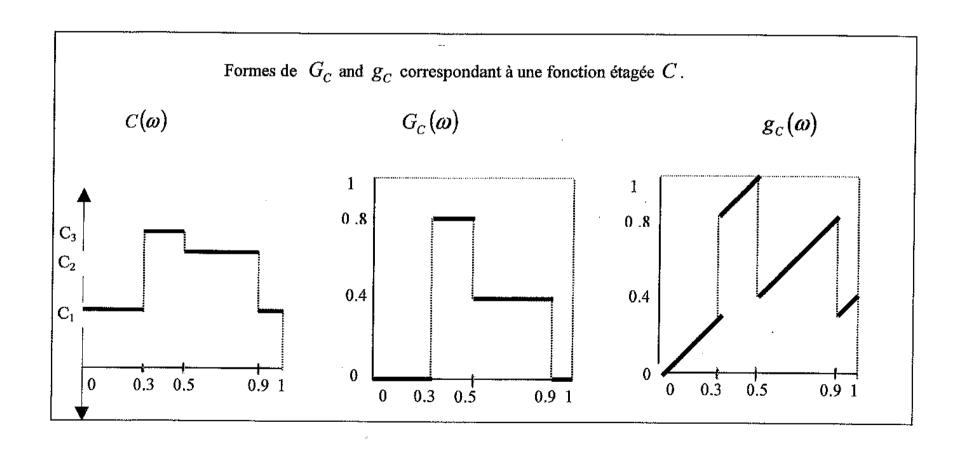
The way to derive g_C from G_C depends on whether or not C has *levels*.

<u>Definition</u>: we call a *level* of criterion C any inverse image of a real number y such that $P^{C}(y) = P(A) > 0$ i.e. C has *levels* when horizontal line segments form part of the graph of C.

Properties of G_C

- The range of G_C is included in [0,1[
- G_C has the same levels as C
- G_C verifies implication (1)
- For every y in the range of G_C , we have $P(u|G_C(u) < y) = y$
- If C has no level, G_C is a coordination function.

If C has at least one level, the range of function G_C is strictly included in [0,1[. We have to deduce from G_C another function, denoted g_C , such that the range of g_C is equal to [0,1[.



Coordination function with several criteria

If criterion C has levels A_i, we can introduce secondary criteria C_i corresponding to each level.

The coordination function verifies the following conditions:

$$\forall \omega_1, \omega_2 \in A_i \ C_i(\omega_1) < C_i(\omega_2) \Rightarrow g(\omega_1) < g(\omega_2)$$

Application to Poisson sampling

$$\underline{Initialization}: \Gamma_{k,0}(\omega_k) = 0 \quad g_{k,1}(\omega_k) = \omega_k$$

$$I_{k,1}(\omega_k) = II_{[0,\pi_{k,1}[}(\omega_k) \quad \Gamma_{k,1}(\omega_k) = \gamma_{k,1} II_{[0,\pi_{k,1}[}(\omega_k)$$

For sample S_t , we choose a coordination function $g_{k,t}$ associated to each unit k. Then:

$$k \in S_t \iff g_{k,t}(\omega_k) \in [0, \pi_{k,t}[$$

We define:
$$A_{k,t} = g_{k,t}^{-1}[0, \pi_{k,t}[$$

$$\rightarrow$$
 indicator function: $I_{k,t}(\omega_k) = \mathbb{I}_{A_{k,t}}(\omega_k)$

Application to Poisson sampling

Sampling

For a separate sample t, for the selection of the sample S_t:

criterion $C_{k,t}(\omega_k)$ = cumulative burden $\Gamma_{k,t-1}(\omega_k)$, and then deduce $g_{k,t}$

For <u>updating a panel</u>: $S_u = \text{sample corresponding to the latest update}$ ($u \le t - 1$). We denote $A_{k,u} = g_{k,u}^{-1} [0, \pi_{k,u}[$

First-stage criterion in the calculation of the coordination function : any decreasing function of the indicator function of $A_{k,u}$.

For example:
$$C_{k,t}(\omega_k) = 1$$
 if $\omega_k \in A_{k,u}$ (i.e. $k \in S_u$)
= 2 if $\omega_k \notin A_{k,u}$ (i.e. $k \notin S_u$)

To take into account past burdens:

cumulative burden $\Gamma_{k,t-1}$ = secondary criterion This leads to a certain coordination function $g_{k,t}$.

Application to SSRS

1. Calculation of the approximate $I'_{k,t}(\omega_k)$, by its conditional expectation

$$I'_{k,t}(\omega_k) = E(I_{k,t}(\Omega) | \Omega = \Omega_k)) = b_{k,t}(g_{k,t}(\omega_k))$$

where $\Omega = (\Omega_1 \dots \Omega_k \dots \Omega_N)$ is a random vector from which we have a realization ω .

2. Approximation by step functions

 $I'_{k,t}$ and the cumulative burden function are no longer step functions. These functions are approximated by step functions, which are constant over predefined intervals, obtained by dividing [0, 1[in L equal subintervals.

Then, the procedures are similar to those used in Poisson sampling.

Simulation

Fixed population of size 100.

Initial burden = 0 for each unit.

Selection of 20 samples (simple random sampling).

For each sample : sample size = 25, response burden = 1,

except:

Samples n°3 and n°15: sample size = 50, response burden = 3

Samples $n^{\circ}10$ and $n^{\circ}11$: sample size =10, response burden = 2

 \rightarrow mean cumulative burden = 7.4

